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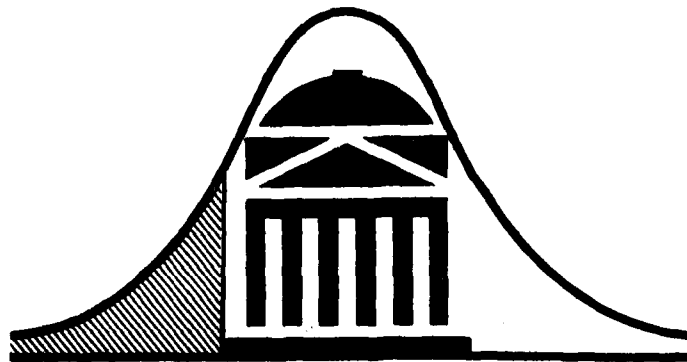
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by

Youn-Min Chou and D. B. Owen

Technical Report No. 147 ✓
Department of Statistics ONR Contract

August, 1981

Research sponsored by the Office of Naval Research
Contract N00014-76-0613

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BIVARIATE CUMULANTS OF A SINGLY TRUNCATED BIVARIATE NORMAL DISTRIBUTION

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SUMMARY

A method of obtaining the bivariate cumulants of any order is given for a truncated bivariate normal distribution where one of the variates is truncated at w_0 . Some representative values are displayed in tables.

Some key words: Bivariate cumulant; Singly truncated bivariate normal distribution.

1. INTRODUCTION

Let the joint density of a standardized bivariate normal distribution be given by

$$\phi(x, y; \rho) = (2\pi)^{-1} (1 - \rho^2)^{-\frac{1}{2}} \exp\{-\frac{1}{2}(x^2 - 2\rho xy + y^2)/(1 - \rho^2)\}$$

for $-\infty < x, y < +\infty$.

We will also use the notation

$$G'(x) = (2\pi)^{-\frac{1}{2}} \exp\{-\frac{1}{2}x^2\},$$

and
$$G(x) = \int_{-\infty}^x G'(t) dt$$

for $-\infty < x < +\infty$,

for the standardized univariate density and cumulative, respectively.

Then if the X variate is truncated below w_0 , the joint density of the singly truncated bivariate normal distribution (STBVND) is given by

$$f(x,y;\rho) = \phi(x,y;\rho)/G(-w_0) \quad \text{for } w_0 < x < +\infty, -\infty < y < +\infty.$$

Our purpose is to obtain the bivariate cumulants corresponding to $f(x,y;\rho)$.

2. BIVARIATE CUMULANTS

Cook (1951) illustrated three methods of deriving bivariate cumulants. Cumulants of all orders of the bivariate distribution may be worked out by choosing the appropriate operation. She gave all the formulae for bivariate cumulants, κ_{ij} , up to $i + j = 6$. As the order of the cumulants increases, the number of terms increases greatly. Johnson and Kotz (1972) gave bivariate cumulants only up to $i + j = 2$. Gajjar and Subrahmaniam (1978) obtained bivariate moments up to order 4. We give here a general formula for bivariate cumulants for any order.

The moment generating function of a STBVND is given by

$$M(t_1, t_2) = [G(-w_0)]^{-1} G(t_1 + \rho t_2 - w_0) \exp\left[\frac{1}{2}(t_1^2 + 2\rho t_1 t_2 + t_2^2)\right].$$

where t_1 corresponds to x and t_2 corresponds to y . Since the cumulant generating function is $\ln[M(t_1, t_2)]$, we have the following expression for the cumulant generating function $K(t_1, t_2)$

$$K(t_1, t_2) = -\ln[G(-w_0)] + \ln[G(t_1 + \rho t_2 - w_0)] + t_1^2/2 + \rho t_1 t_2 + t_2^2/2$$

and the cumulant κ_{ij} is obtained by taking the i -th partial derivative with respect to t_1 and the j -th partial derivative with respect to t_2 and setting $t_1 = t_2 = 0$.

We obtain

$$\kappa_{10} = G'(w_0)/[G(-w_0)]$$

$$\kappa_{01} = \rho \kappa_{10}$$

$$\kappa_{20} = w_0 G'(w_0)/[G(-w_0)] - [G'(w_0)/G(-w_0)]^2 + 1$$

$$= w_0 \kappa_{10} - \kappa_{10}^2 + 1$$

$$\kappa_{02} = \rho^2 (\kappa_{20} - 1) + 1$$

and

$$\kappa_{ij} = \rho^j \left. \frac{\partial^{i+j-1}}{\partial x^{i+j-1}} \left(\frac{G'(x)}{G(x)} \right) \right|_{x = -w_0} \quad \text{for } i + j \neq 2$$

The expressions for $\frac{\partial^i}{\partial x^i} \left(\frac{G'(x)}{G(x)} \right)$ become very cumbersome as is

illustrated by the following.

$$\frac{\partial}{\partial x} \left(\frac{G'(x)}{G(x)} \right) = -x \left(\frac{G'(x)}{G(x)} \right) - \left(\frac{G'(x)}{G(x)} \right)^2,$$

$$\frac{\partial^2}{\partial x^2} \left(\frac{G'(x)}{G(x)} \right) = (x^2 - 1) \left(\frac{G'(x)}{G(x)} \right) + 3x \left(\frac{G'(x)}{G(x)} \right)^2 + 2 \left(\frac{G'(x)}{G(x)} \right)^3,$$

$$\begin{aligned} \frac{\partial^3}{\partial x^3} \left(\frac{G'(x)}{G(x)} \right) &= (-x^3 + 3x) \left(\frac{G'(x)}{G(x)} \right) - (7x^2 - 4) \left(\frac{G'(x)}{G(x)} \right)^2 \\ &\quad - 12x \left(\frac{G'(x)}{G(x)} \right)^3 - 6 \left(\frac{G'(x)}{G(x)} \right)^4. \end{aligned}$$

Hence, we look for a way to generate these derivatives recursively.

We let $g = g(x) = G'(x)/[G(x)]$ and $h = 1/g = G(x)/[G'(x)]$ and note

that the derivatives of h are easily obtained as

$$h' = 1 + xh$$

$$h^{(n)} = (n-1)h^{(n-2)} + xh^{(n-1)} \quad \text{for } n \geq 2, \text{ where } h^{(0)} \equiv h.$$

Since $gh = 1$, an application of Liebniz's rule gives

$$\sum_{j=0}^n \binom{n}{j} g^{(n-j)} h^{(j)} = 0$$

or

$$g^{(n)} = -g \sum_{j=1}^n \binom{n}{j} g^{(n-j)} h^{(j)}.$$

Hence, we can obtain any derivative of $G'(x)/G(x)$ in terms of lower order derivatives and derivatives of h which are given by the recursion formula for $h^{(n)}$.

3. APPLICATION

In the problem of screening based on a singly truncated bivariate normal distribution, one needs to know the distribution of the sample correlation coefficient. This can be achieved by supplying the bivariate cumulants given above to Gayen's (1951) results.

4. TABLE

In the accompanying table we give values of κ_{10} , κ_{20} , κ_{30} , κ_{40} and κ_{50} . These cumulants are independent of ρ . However, the remaining cumulants up to order 5 may be obtained from these, using the following simple formulas which were obtained from the general formula for κ_{ij} given above:

$$\kappa_{01} = \rho \kappa_{10}$$

$$\kappa_{02} = \rho^2 (\kappa_{20} - 1) + 1$$

$$\kappa_{11} = \rho \kappa_{20}$$

$$\kappa_{ij} = \rho^j \kappa_{i+j,0} \quad \text{for } i+j > 2$$

TABLE 1

BIVARIATE CUMULANTS OF A SINGLY TRUNCATED
BIVARIATE NORMAL DISTRIBUTION

w_0	κ_{10}	κ_{20}	κ_{30}	κ_{40}	κ_{50}
-3.0	.004438	.986667	.035680	-.081046	.139672
-2.8	.007936	.977717	.054810	-.110766	.154800
-2.6	.013647	.964333	.080062	-.141556	.148857
-2.4	.022580	.945299	.111173	-.168418	.114778
-2.2	.035975	.919561	.146778	-.185535	.052033
-2.0	.055248	.886452	.184395	-.187855	-.031092
-1.8	.081893	.845887	.220752	-.172786	-.118818
-1.6	.117352	.798466	.252404	-.141250	-.192852
-1.4	.162881	.745436	.276436	-.097540	-.238752
-1.2	.219437	.688524	.291034	-.048050	-.250453
-1.0	.287600	.629686	.295718	.000547	-.230966
-.8	.367562	.570849	.291238	.042875	-.189622
-.6	.459147	.513695	.279206	.075707	-.137889
-.4	.561883	.459534	.261660	.098016	-.085824
-.2	.675073	.409261	.240659	.110477	-.040233
.0	.797885	.363380	.218015	.114769	-.004433
.2	.929416	.322069	.195158	.112946	.020990
.4	1.068757	.285262	.173110	.106994	.037137
.6	1.215026	.252727	.152523	.098587	.045869
.8	1.367403	.224132	.133752	.089010	.049170
1.0	1.525136	.199096	.116935	.079165	.048804
1.2	1.687553	.177229	.102061	.069643	.046173
1.4	1.854058	.158150	.089030	.060786	.042308
1.6	2.024130	.141506	.077687	.052765	.037920
1.8	2.197314	.126976	.067859	.045637	.033466
2.0	2.373217	.114276	.059367	.039387	.029222
2.2	2.551498	.103155	.052040	.033956	.025338
2.4	2.731863	.093396	.045722	.029266	.021881
2.6	2.914059	.084813	.040273	.025234	.018866
2.8	3.097868	.077244	.035568	.021774	.016280
3.0	3.283101	.070551	.031501	.018808	.014096

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1. REPORT NUMBER 147	2. GOVT ACCESSION NO. AD-A104914	3. RECIPIENT'S CATALOG NUMBER
4. TITLE (and Subtitle) BIVARIATE CUMULANTS OF A SINGLY TRUNCATED BIVARIATE NORMAL DISTRIBUTION		5. TYPE OF REPORT & PERIOD COVERED 11 TECHNICAL REPORT
7. AUTHOR(s) Youn-Min Chou and D. B. Owen		6. PERFORMING ORG. REPORT NUMBER 147
9. PERFORMING ORGANIZATION NAME AND ADDRESS Southern Methodist University Dallas, Texas 75275		8. CONTRACT OR GRANT NUMBER(s) 14 N00014-76-C-0613
11. CONTROLLING OFFICE NAME AND ADDRESS Office of Naval Research Arlington, VA 22217		10. PROGRAM ELEMENT, PROJECT, TASK AREA & WORK UNIT NUMBERS NR 042-389
14. MONITORING AGENCY NAME & ADDRESS (if different from Controlling Office) 1011		12. REPORT DATE August 1981
		13. NUMBER OF PAGES 6
		15. SECURITY CLASS. (of this report)
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18. SUPPLEMENTARY NOTES		
19. KEY WORDS (Continue on reverse side if necessary and identify by block number) Bivariate cumulant; Singly truncated bivariate normal distribution.		
20. ABSTRACT (Continue on reverse side if necessary and identify by block number) A method of obtaining the bivariate cumulants of any order is given for a truncated bivariate normal distribution where one of the variates is truncated at w_0 . Some representative values are displayed in tables.		